

Exports and Credit Constraints under Incomplete Information: Theory and Evidence from China

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- The financial crisis of 2008 raises the question of whether credit constraints faced by exporters played a role in the fall in world trade.
- Empirical findings on this question are divergent: Amiti and Weinstein (2009) and Chor and Manova (2010) vs. Levchenko et. al (2010) and Behrens, Corcos and Mion (2010).
- This paper go back to the theory and ask why credit for exports should be allocated any differently than credit for domestic sales. We emphasize one possible reason, longer time-lag between production and export sales for exports than for domestic sales.
- This reason is built into a model of heterogeneous firms with incomplete information.

Introduction continued

- Assumptions: a single bank makes loans of working capital, and has incomplete information about the firms in two respects:
 - Does not know whether a loan is used for domestic or export production: this means we are not modeling "export finance"
 - Does not know the productivity of firms: this means that without loss of generality we can focus on incentive-compatible loan policies
- Goals of the paper:
 - We solve for the incentive-compatible loan policy for the bank, building in the longer time-lag for exporters
 - We estimate a structural equation arising from the model, under which sales revenue is a function of interest payments and other variables, on a dataset of Chinese domestic and exporting firms over 2000 - 2008.

- We obtain robust empirical evidence that exporting firms face more severe credit constraints than purely domestic firms.
The credit constraint is more stringent
 - as a firm's export share grows,
 - as the time to ship for exports is lengthened
 - and as there is greater dispersion of firms' productivities reflecting information incompleteness.

- Monopolistic competition with heterogeneous firms, productivity x
- Consumer has CES utility function and demand system.
- Firms need to cover a fraction δ of total costs, taking loan $M(x')$ and interest payment $I(x')$ schedule as given, and choosing quantity q and claimed productivity x'

- Domestic firm maximizes:

$$\begin{aligned} \max_{x', q_d} \pi_d(x, x') &= p_d q_d - (1 - \delta) \left(\frac{q_d w}{x} + C_d \right) - (M_d(x') + I_d(x')), & (1) \\ \text{s.t. } \pi_d(x, x) &\geq \pi_d(x, x'), \\ \pi_d(x, x) &\geq 0, \\ M_d(x') &\geq \delta \left(\frac{q_d w}{x} + C_d \right), \end{aligned}$$

and also CES demand.

- First constraint is the incentive compatibility condition.
- Second constraint is zero-cutoff-profits
- Third constraint is for the loan

- The incentive compatibility condition implies locally:

$$\left. \frac{\partial \pi_d(x, x')}{\partial x'} \right|_{x'=x} = 0.$$

- Use the loan constraint to solve for quantity produced, and therefore price and revenue from the CES demand system. The above condition becomes:

$$[\Phi_d(x, M_d(x)) - 1] \frac{M'_d(x)}{\delta} = I'_d(x),$$

where

$$\Phi_d(x, M_d(x)) \equiv \left[p_d \left(\frac{\sigma - 1}{\sigma} \right) \right] / \frac{w}{x},$$

- Φ_d is the ratio of expected marginal revenue to marginal costs.
- Without any credit constraint, expected marginal revenue equals marginal cost, so that $\Phi_d = 1$.
- When firms produce quantity below the optimal level then $\Phi_d > 1$, indicating a credit constraint.

- The intuition of the credit constraints:
 - Suppose that $M'_d(x)$ and $I'_d(x)$ are both positive \Rightarrow the firm *must be* credit constrained, i.e. $\Phi_d > 1$.
 - A firm that is producing at the first-best would have only a *second-order loss* in profits from announcing a slightly smaller productivity x' , and producing at a slightly lower level.
 - But the firm would have a *first-order gain* from the reduction in interest payments, with $I'_d(x) > 0$.
 - So a firm at the first-best would always understate its productivity.
 - With increasing loan and interest payment schedules, a credit constraint is therefore needed to ensure incentive compatibility.

- Exporters' decision:

$$\begin{aligned} \max_{x', q_d, q_e} \pi_e(x, x') &= p_d q_d + p_e q_e - (1 - \delta) \left(\frac{q_d \bar{w}}{x} + C_d + \frac{q_e \bar{w}}{x} + C_e \right) \\ &\quad - (M_e(x') + I_e(x')) , \\ \text{s.t. } \pi_e(x, x) &\geq \pi_e(x, x') \\ \pi_e(x, x) &\geq \pi_d(x, x) \\ M_e(x') &\geq \delta \left(\frac{q_d \bar{w}}{x} + C_d + \frac{q_e \bar{w}}{x} + C_e \right) , \end{aligned} \quad (2)$$

and CES demand for exports.

- The bank is making a single loan $M_e(x')$ and cannot verify whether the loan is used to cover the costs of production for domestic sales or for exports.

Exporting Firm continued

- Exporters' quantity decision:

$$p_d \left(\frac{\sigma - 1}{\sigma} \right) = p_e \left(\frac{\sigma - 1}{\sigma} \right). \quad (3)$$

For any given loan, the bank will know exactly how production is allocated between the two market. We break up the total loan into the domestic component, $M_e^d(x')$, and the export component, $M_e^e(x')$.

- Exporters' incentive-compatibility condition:

$$\left[\Phi_e^d(x, M_e^d(x)) - 1 \right] \frac{M_e^{d'}(x)}{\delta} + \left[\Phi_e^e(x, M_e^e(x)) - 1 \right] \frac{M_e^{e'}(x)}{\delta} = I_e'(x), \quad (4)$$

where,

$$\Phi_e^d(x, M_e^d(x)) \equiv \left[p_d \left(\frac{\sigma - 1}{\sigma} \right) \right] / \frac{w}{x} \quad (5)$$

$$\Phi_e^e(x, M_e^e(x)) \equiv \left[p_e \left(\frac{\sigma - 1}{\sigma} \right) \right] / \frac{w}{x}$$

$$\Phi_e^d(x, M_e^d(x)) = \Phi_e^e(x, M_e^e(x)).$$

- The bank faces an opportunity cost of i – the interest rate – on its loans.
- We assume that the loans for domestic (export) projects are paid back after τ_d (τ_e) periods, and further assume that $\tau_e > \tau_d$, reflecting the longer time-lags involving in the shipping of exports.
- Export loans therefore have a higher opportunity cost for the bank, which is a potential reason to treat exporters and domestic firms differently.

The Bank continued

- Bank's decision:

$$\max_{M, I} \int_{\underline{x}_d}^{\underline{x}_e} (I_d(x) - i\tau_d M_d(x)) f(x) dx + \int_{\underline{x}_e}^{\infty} (I_e(x) - i\tau_d M_e^d(x) - i\tau_e M_e^e(x)) f(x) dx \quad (6)$$

s. t.

$$[\Phi_d(x, M_d(x)) - 1] \frac{M_d'(x)}{\delta} = I_d'(x), \text{ if } x \in [\underline{x}_d, \underline{x}_e)$$

and

$$[\Phi_e^d(x, M_e^d(x)) - 1] \frac{M_e^{d'}(x)}{\delta} + [\Phi_e^e(x, M_e^e(x)) - 1] \frac{M_e^{e'}(x)}{\delta} = I_e'(x), \text{ if } x \in [\underline{x}_e, \infty)$$

- First solve optimal control problem, and then solve for optimal \underline{x}_d and \underline{x}_e

- Assuming Pareto distribution of productivity, solution to the optimal control problem is:

$$\Phi_d(x, M_d(x)) = \bar{\Phi}_d \equiv (1 + i\delta\tau_d) \left(1 - \frac{\sigma-1}{\sigma\theta}\right)^{-1}, \quad (7)$$

$$\Phi_e^d(x, M_e^d(x)) = \Phi_e^e(x, M_e^e(x)) = \bar{\Phi}_e \equiv [1 + i\delta(\tau_d\eta_d + \tau_e\eta_e)] \left(1 - \frac{\sigma-1}{\sigma\theta}\right)^{-1}.$$

where η_d and η_e are the shares of demand coming from the domestic and foreign markets as:

$$\eta_d = \frac{\gamma P^{\sigma-1}}{\gamma P^{\sigma-1} + \gamma^* P^{*\sigma-1}} \text{ and } \eta_e = \frac{\gamma^* P^{*\sigma-1}}{\gamma P^{\sigma-1} + \gamma^* P^{*\sigma-1}}.$$

- (a) credit constraint applies even if $i = 0$ (like in earlier intuition);
- (b) exporters face a tighter credit constraint in domestic market than domestic firms, $\bar{\Phi}_e > \bar{\Phi}_d$ when $\tau_e > \tau_d$ and $i > 0$ (opportunity cost to bank);
- (c) exporters face the same credit constraint in export market and domestic market, since $\Phi_e^d = \Phi_e^e = \bar{\Phi}_e$

Cutoff Productivities

- The bank will optimally choose the initial interest payment for a domestic firm to determine \underline{x}_d via the zero-cutoff-profit condition $\pi_d(\underline{x}_d, \underline{x}_d) = 0$
- Likewise the initial interest payment for the cutoff exporter is chosen to determine \underline{x}_e via $\pi_d(\underline{x}_e, \underline{x}_e) = \pi_e(\underline{x}_e, \underline{x}_e)$
- The solution for the cutoff firms are:

$$\frac{M_d(\underline{x}_d)}{\delta} = \sigma C_d,$$
$$\frac{M_e(\underline{x}_e)}{\delta} = (\Delta(\sigma - 1) + 1) C_e + C_d,$$

where

$$\Delta \equiv \left(\frac{1 + i\delta\tau_e}{1 + i\delta(\tau_d\eta_d + \tau_e\eta)} \right) \left(1 - \left(\frac{1 + i\delta(\tau_d\eta_d + \tau_e\eta_e)}{1 + i\delta\tau_d} \right)^{\sigma-1} \eta_d \right)^{-1},$$
$$\Theta \equiv \frac{i\delta(\tau_e - \tau_d)}{\left(1 - \frac{\sigma-1}{\sigma\theta}\right)} (\eta_d C_e - \eta_e C_d).$$

- The credit constraints have impacts on the extensive margin:
 - σC_d is the total cost of the credit unconstrained production for the cutoff producer in Melitz (2003). Here, since there are credit constraints so $\bar{\Phi}_d > 1$, the cutoff productivity \underline{x}_d is *higher* than the cutoff productivity in Melitz (2003).
 - When $i = 0$ so that $\Delta = 1/\eta_e$, then $\frac{M_e(\underline{x}_e)}{\delta}$ is the export costs of production for the cutoff exporter in Melitz (2003). Likewise, since $\bar{\Phi}_e > 1$, the cutoff exporter \underline{x}_e is also higher than that in Melitz (2003).

- We derive a linear relationship between interest payments and revenue of the firm

$$r_{jt}(x) = \beta_0 C + \beta_1 I_{jt}(x) + g_1(\eta_e) I_{jt}(x) + g_2(\eta_e) C + g_3(\eta_e) i \mathbf{1}_{\{x_{jt} \geq x_e\}}, \quad (8)$$

where

$$I(x) = \begin{cases} I_d(x) & \text{if } x \in [\underline{x}_d, \underline{x}_e] \\ I_e(x) & \text{if } x \in [\underline{x}_e, \infty] \end{cases}, \quad r(x) = \begin{cases} p_d q_d & \text{if } x \in [\underline{x}_d, \underline{x}_e] \\ p_d q_d + p_e q_e & \text{if } x \in [\underline{x}_e, \infty] \end{cases}$$

and $\mathbf{1}_{\{x_{jt} \geq x_e\}}$ is an indicator variable which takes one for $x_{jt} \geq x_e$ and zero otherwise;

$$C = \begin{cases} C_d & \text{if } x \in [\underline{x}_d, \underline{x}_e] \\ C_d + C_e & \text{if } x \in [\underline{x}_e, \infty] \end{cases}$$

- The coefficients are obtained from the model as:

$$\begin{aligned}\beta_0 &= -\frac{\sigma}{\sigma-1}\bar{\Phi}_d < 0, \\ \beta_1 &= \frac{\sigma}{\sigma-1}\left(\frac{\bar{\Phi}_d}{\bar{\Phi}_d-1}\right) > 0,\end{aligned}\tag{9}$$

- Fixed costs lower the loan available to cover variable cost $\beta_0 < 0$
- a larger payments are associated with larger revenue $\beta_1 > 0$
- $g_1(\eta_e), g_2(\eta_e), g_3(\eta_e)$ are functions of export share η_e

Estimating Equations continued

- The coefficients are obtained from the model as:

$$g_1(\eta_e) = \frac{\sigma}{\sigma-1} \left(\frac{\bar{\Phi}_e}{\bar{\Phi}_e-1} - \frac{\bar{\Phi}_d}{\bar{\Phi}_d-1} \right) \leq 0,$$

$$g_2(\eta_e) = -\frac{\sigma}{\sigma-1} (\bar{\Phi}_e - \bar{\Phi}_d) \leq 0,$$

$$g_3(\eta_e) = -\frac{\sigma}{\sigma-1} \left(\frac{\bar{\Phi}_e}{\bar{\Phi}_e-1} \right) \frac{\delta(\tau_e - \tau_d)}{(1 - \frac{\sigma-1}{\sigma\theta})} (\eta_d C_e - \eta_e C_d).$$

- $g_j(0) = 0$ for $j = 1, 2$, and $g_j(\eta_e) < 0$ for $j = 1, 2$, $\eta_e > 0$, and $\tau_e > \tau_d$ and interest rate $i > 0$, so that $\bar{\Phi}_e > \bar{\Phi}_d$.
- For exporters, bank payments are positively associated with revenue of $0 < \beta_1 + g_1(\eta_e) < \beta_1$.
- Similarly, $0 > \beta_0 + g_2(\eta_e) > \beta_0$, indicating fixed costs reduce exporter's revenue, but less than β_0 .
- So exporters are constrained in what they can earn due to the extra credit constraint that they face via both exporter's bank payments and their fixed costs.
- In addition, exporters face a reduction in revenue from any increase in the interest rate i , as shown by the final term $g_3(\eta_e)$, conditional on the exporting fixed costs or domestic market share being large enough so that $\eta_d C_e - \eta_e C_d > 0$.

Estimating Equations continued

- To estimate the coefficients, we simplify them in the following ways:

$$\begin{aligned}g_1(\eta_e) &= -\frac{\sigma}{\sigma-1} \frac{i\delta\eta_e(\tau_e - \tau_d)}{(i\delta(\tau_d\eta_d + \tau_e\eta_e) + \frac{\sigma-1}{\sigma\theta})(i\delta\tau_d + \frac{\sigma-1}{\sigma\theta})} \\&\simeq -\frac{\sigma}{\sigma-1} \frac{1}{(i\delta\tau_d + \frac{\sigma-1}{\sigma\theta})} \left(\left(\frac{i\delta(\tau_e - \tau_d)}{i\delta\tau_d + \frac{\sigma-1}{\sigma\theta}} \right) \eta_e - \left(\frac{i\delta(\tau_e - \tau_d)}{i\delta\tau_d + \frac{\sigma-1}{\sigma\theta}} \right)^2 \eta_e^2 \right) \\&= \beta_2\eta_e + \beta_3\eta_e^2.\end{aligned}$$

- where $\beta_2 = -\frac{\sigma}{\sigma-1} \frac{1}{(i\delta\tau_d + \frac{\sigma-1}{\sigma\theta})} \left(\frac{i\delta(\tau_e - \tau_d)}{i\delta\tau_d + \frac{\sigma-1}{\sigma\theta}} \right) < 0$ and

$$\beta_3 = \frac{\sigma}{\sigma-1} \frac{1}{(i\delta\tau_d + \frac{\sigma-1}{\sigma\theta})} \left(\frac{i\delta(\tau_e - \tau_d)}{i\delta\tau_d + \frac{\sigma-1}{\sigma\theta}} \right)^2 > 0.$$

- So

$$g_1(\eta_e) = -\frac{\beta_2^2}{\beta_3} \left(\frac{1}{1 - (\beta_2/\beta_3)\eta_e} \right).$$

Estimating Equations continued

- The property of $g_1(\eta_e) = -\frac{\beta_2^2}{\beta_3} \left(\frac{1}{1-(\beta_2/\beta_3\eta_e)} \right)$:
 - $\beta_2 < 0$ and $\beta_3 > 0$ ensure that $g_1(\eta_e) < 0$ for $\eta_e > 0$.
 - the formula $g_1(\eta_e)$ shows that $|g_1(\eta_e)| < \beta_1$, which ensures there is still a positive relationship between bank payments and revenue for exporters.
 - Provided that $\beta_2 < 0$ and $\beta_3 > 0$, $g_1(\eta_e)$ is decreasing in the export share, which means that exporting firms face more credit constraints if their export share is higher.
 - Using $\eta_e = 1$, $\beta_2 < 0$ and $\beta_3 > 0$, we get $|g_1(1)| < \beta_1$ holds if and only if

$$\beta_2 \in \left(-\frac{1}{2}(\beta_1 + \sqrt{\beta_1^2 + 4\beta_1\beta_3}), 0 \right). \quad (10)$$

- Turning to the function g_2 ,

$$g_2(\eta_e) = -\frac{\sigma}{\sigma-1} i \delta \eta_e (\tau_e - \tau_d) \left(1 - \frac{\sigma-1}{\sigma\theta} \right)^{-1} \quad (11)$$

$$\equiv \beta_4 \eta_e, \quad (12)$$

Estimating Equations continued

- The estimation equation is:

$$R_{jt} = \beta_0 + (\beta_1 + \beta_2\eta_e + \beta_3\eta_e^2) I_{jt}(x) + \beta_4\eta_e + \beta_5\mathbf{1}_{\{x_{jt} \geq \bar{x}_e\}} + \varepsilon_{jt}. \quad (13)$$

- Instruments:

- we need variables that are independent of the error ε_{jt} , reflecting all *ex-post* shocks to production.
- We use the *ex-ante* productivity that is anticipated by firms as the instrument.
- Using the the technique of Olley and Pakes (1996), we get
 - total factor productivity (TFP) of the firm *inclusive* of the unanticipated, random productivity shocks (what we call TFP1). This is the standard firm-level measure of productivity.
 - TFP of the firms *exclusive* of these unanticipated shocks (what we call TFP2). This makes use of the firm investment decisions to infer the productivity that is anticipated by the firm.
- By construction, TFP2 is independent of the unanticipated shocks within the error ε_{jt} .

- A Chinese firm-level panel data set which covers more than 160,000 manufacturing firms per year over 2000-2008.
- The dataset covers two types of manufacturing firms:
 - All state-owned enterprises (SOEs) – few and omitted here;
 - Non-SOEs whose annual sales are more than five million *Renminbi* (which is equivalent to around \$735,000 under current exchange rate).
- The data set includes many financial variables listed in the main accounting sheets of all these firms.
- Another Chinese transaction-level trade data set for the years 2000-2006, including shipments and export value for each exported good.

Benchmark estimates

Table: Benchmark Estimates for Chinese and Foreign Firms (2000-2008)

Data Sample	For Chinese-owned firms			Foreign firms
Regressand: Firm's Revenue	OLS	2SLS	2SLS	2SLS
	(1)	(2)	(3)	(4)
Interest Payment (β_1)	64.46*** (31.18)	82.44*** (32.23)	84.99*** (32.38)	33.97*** (3.13)
Interest Payment × Export Share (β_2)	-67.55** (-2.43)	-112.3** (-2.00)	-212.5*** (-3.94)	242.9* (1.93)
Interest Payment × Export Share Squared (β_3)	165.1*** (4.43)	592.9*** (2.70)	910.2*** (4.37)	390.6 (1.31)
Export Share (β_4)	-12,585*** (-9.15)	-60,334*** (-14.52)	-64,818*** (-11.92)	49,957*** (2.59)
Export Indicator (β_5)	7,693*** (7.45)	-304.6 (-1.03)	-936.4*** (-2.59)	-1,922* (-1.67)
Lower Bound $-\frac{1}{2}(\beta_1 + \sqrt{\beta_1^2 + 4\beta_1\beta_3})$	-140.3	-266.1	-323.8	-133.4
Mean of Fitted Export Share (η_e^m)	0.096	0.095	0.090	0.326
Estimated Value of $g_1(\eta_e^m)$	-5.25	-7.10	-13.80	166.4
90% of Fitted Export Share (η_e^u)	0.462	0.210	0.183	0.501
Estimated Value of $g_1(\eta_e^u)$	-14.65	-11.18	-21.79	626.1
Industry-Specific Fixed Effects	Yes	Yes	Yes	Yes
Year-Specific Fixed Effects	Yes	Yes	Yes	Yes
Hong Kong/Taiwan/Macao Firms Included	Yes	Yes	No	No
Observations	926,534	909,229	812,558	99,886
First-Stage Regressions				
Kleibergen-Paap rk LM χ^2 statistic		367.4 [†]	30.06 [†]	51.05 [†]
Kleibergen-Paap rk Wald F statistic		222.5 [†]	8.11 [†]	41.34 [†]
IV1: $TFP_{2,it}$		-409*** (-3.85) [37.33]	-241*** (-4.58) [16.74]	-2.977*** (-3.51) [15.00]

Bivariate Selection

Table 4: The Heckman Two-Step Estimates of Bivariate Selection Model (2000-2008)

Heckman Two-step: Regressand:	Chinese Firms		Foreign Firms	
	1 st Step	2 nd Step	1 st Step	2 nd Step
	Export Indicator	Export Share	Export Indicator	Export Share
	(1)	(2)	(3)	(4)
Log of TFP	0.011*** (6.66)		-0.015*** (-3.81)	
Export-based EG Index	0.023*** (6.81)	0.023*** (37.87)	0.005 (0.48)	-0.002 (-0.56)
Tangible Assets Ratio	1.034*** (28.06)	1.067*** (62.35)	1.329*** (11.52)	-0.368*** (-5.03)
Intangible Assets Indicator	-0.504*** (-101.10)	-0.519*** (-65.85)	-0.362*** (-28.48)	0.112*** (6.06)
Log of Capital-Labor Ratio	0.007*** (5.62)	-0.006*** (-26.14)	-0.005 (-1.63)	-0.023*** (-24.51)
Inverse Mills Ratio		1.180*** (58.02)		-0.903*** (-10.35)

Sea shipments

Table 5: 2SLS Estimates by Sea and Non-sea Shipments for Chinese Firms

Data Sample:	All Firms (2000-2008)	Matched Firms (2000-2006)		
		All Matched Firms	Interact with Sea Dummy	Interact with Non-sea Dummy
Regressand: Firm's Revenue	(1)	(2)	(3)	(4)
Interest Payment (β_1)	76.85*** (30.96)	86.31*** (23.83)		85.76*** (23.54)
Interest Payment	-303.8***	-539.7***	-638.1***	-511.5***
× Fitted Export Share (β_2)	(-7.75)	(-10.27)	(-7.89)	(-9.93)
Interest Payment	1,211***	1,457***	2,011***	1,329***
× Fitted Export Share Squared (β_3)	(11.95)	(10.84)	(7.05)	(9.88)
Fitted Export Share (β_4)	-236,380*** (-8.76)	-210,621*** (-9.30)	-220,669*** (-9.04)	-207,623*** (-9.24)
Export Indicator (β_5)	2,433*** (3.77)	3,552*** (6.72)		3,767*** (6.48)
Tangible Asset Ratio	66,817*** (10.85)	69,260*** (12.93)		69,201*** (12.87)
Tangible Asset Ratio	-18,745*	-12,801		-16,361*
× Fitted Export Share	(-1.81)	(-1.30)		(-1.67)
Tangible Asset Ratio	12,287	35,897***		40,034***
× Fitted Export Share Squared	(0.75)	(2.79)		(3.08)
Intangible Asset Indicator	-19,240*** (-8.82)	-19,886*** (-10.73)		-19,785*** (-10.71)
Mean of Fitted Export Share (η_e^m)	.090	.114	.135	.094
Estimated Value of $g_1(\eta_e^m)$	-20.61	-41.04	-60.65	-38.90
Year-Specific Fixed Effects	Yes	Yes	Yes	Yes
Industry-Specific Fixed Effects	Yes	Yes	Yes	Yes
Number of Observations	812,558	536,100		536,100

Sectoral Productivity Dispersion

Table 6: 2SLS Estimates with Measures of Sectoral Productivity Dispersion (2000-2008)

Regressand: Firm's Revenue	(1)	(2)	(3)	(4)
Percentile of Sectoral Variance of TFP2	<10%	<25%	[25%, 50%]	>90%
Interest Payment (β_1)	43.05*** (14.01)	42.85*** (11.83)	74.53*** (15.97)	103.8*** (16.46)
Interest Payment × Fitted Export Share (β_2)	-100.3*** (-3.65)	-156.8*** (-5.12)	-523.8*** (-6.67)	-1,500*** (-4.22)
Interest Payment × Fitted Export Share Squared (β_3)	378.6*** (5.81)	493.6*** (7.21)	1,130*** (6.74)	4,395*** (3.14)
Fitted Export Share (β_4)	-176,369*** (-10.74)	-307,393*** (-10.50)	-327,343*** (-12.79)	-888,506*** (-4.38)
Export Indicator (β_5)	2,313*** (7.15)	4,563*** (7.95)	5,4911*** (7.55)	17,779*** (4.44)
Tangible Asset Ratio	66,937*** (16.14)	107,471*** (12.52)	91,390*** (12.10)	203,734*** (5.40)
Tangible Asset Ratio × Fitted Export Share	-88,289*** (-10.02)	-100,355*** (-11.03)	-13,216 (-1.27)	562,492*** (4.09)
Tangible Asset Ratio × Fitted Export Share Squared	141,985*** (11.10)	193,605*** (13.38)	135,371*** (7.69)	-1,217,340*** (-4.70)
Intangible Asset Indicator	-18,386*** (-13.20)	-31,819*** (-11.02)	-29,239*** (-11.82)	-70,985*** (-4.78)
Cutoffs of Sectoral Variance of TFP2	<0.566	<0.670	[0.670, 0.915]	>1.599
Mean of Fitted Export Share (η_e^m)	0.169	0.172	0.101	0.084
Estimated Value of $g_1(\eta_e^m)$	-10.36	-17.54	-43.65	-101.8
Industry-Specific Fixed Effects	Yes	Yes	Yes	Yes
Year-Specific Fixed Effects	Yes	Yes	Yes	Yes
Observations	178,388	235,033	180,551	78,939

- We consider a model of heterogeneous firms needing loans for working capital, provided under incomplete information
- Obtain new results on the need for credit constraints. These are tighter for exporting firms due to the longer time-lag between production and receipt of sales revenue
- Empirical results over 2000-2008 support the tighter credit constraint for exporters